

Bahl, Cocke, Jelinek and Raviv (BCJR) Algorithm

Reference: “Optimal Decoding of Linear Codes for Minimizing Symbol Error Rate”, IEEE Trans. Info. Theory, Vol IT-20, pp 284-287, March 1974.

The Viterbi Algorithm – MLSE – minimizes the probability of sequence (word) error

- not necessarily minimize the probability of bit(symbol) error.

Objective: To derive an optimal decoding method for linear codes which minimizes the symbol error probability.

General Problem:

Estimating the *a posteriori* probabilities (APP) of

1. The states and
2. The transitions of a

Markov source observed through a noisy discrete memoryless channel (DMC).



The source above is assumed to be a “discrete-time finite-state” Markov process (e.g. Convolutional encoder, or a finite state machine – FSM).

M states, $m=0,1,\dots,M-1$.

The state of the source at time t is denoted by S_t and its output by X_t .

A state sequence from t to t'

$$S_t^{t'} = S_t, S_{t+1}, \dots, S_{t'}$$

The output sequence $X_t^{t'} = X_t, X_{t+1}, \dots, X_{t'}$.

The state transitions of the Markov source are governed by the transition probabilities

$$p_t(m | m') = \Pr\{S_t = m | S_{t-1} = m'\}$$

And the output by the probabilities

$$q_t(X | m', m) = \Pr\{X_t = X | S_{t-1} = m' ; S_t = m\}$$

where X belongs to some finite discrete alphabet.

The Markov source starts in the initial state $S_0 = 0$, and produces an output sequence X_1^τ ending in the terminal state $S_\tau = 0$.

X_1^τ is the input to a noisy DMC whose output is the sequence

$$Y_1^\tau = Y_1, Y_2, \dots, Y_\tau.$$

The transition probabilities of the DMC are defined by $R(\cdot|\cdot)$ so that for all

$$1 \leq t \leq \tau \quad \Pr\{Y_1^t | X_1^t\} = \prod_{j=1}^t R(Y_j | X_j)$$

The objective of the decoder is to examine Y_1^τ and estimate the APP of the states and the transitions of the Markov source, i.e., the conditional probabilities

$$\Pr\{S_t = m | Y_1^\tau\} = \Pr\{S_t = m; Y_1^\tau\} / \Pr\{Y_1^\tau\} \quad (1)$$

$$\Pr\{S_{t-1} = m'; S_t = m | Y_1^\tau\} = \Pr\{S_{t-1} = m'; S_t = m; Y_1^\tau\} / \Pr\{Y_1^\tau\} \quad (2)$$

Graphical interpretation:

State diagram: Nodes are states, branches – the transitions having nonzero probabilities.

Trellis diagram: Index the states with both the time index t and state index $m \Rightarrow$ “trellis”

Shows the time progression of the state sequences.

For every state sequence S_1^τ there is a unique path through the trellis and vice versa.

If the Markov source is time variant, then we can no longer represent it by a state-transition diagram; however we can construct a trellis for its state sequences.

Associated with each node in the trellis is the corresponding APP

$$\Pr \{ S_t = m \mid Y_1^\tau \} \text{ and}$$

Associated with each branch in the trellis is the corresponding APP

$$\Pr \{ S_{t-1} = m'; S_t = m \mid Y_1^\tau \}.$$

The objective of the decoder is to examine Y_1^τ and compute these APP.

It is easier to derive the joint probabilities

$$\lambda_t(m) = \Pr \{ S_t = m ; Y_1^\tau \} \quad \text{and}$$

$$\sigma_t(m',m) = \Pr \{ S_{t-1} = m'; S_t = m ; Y_1^\tau \}.$$

For a given Y_1^τ , $\Pr \{ Y_1^\tau \}$ is a constant, we can divide $\lambda_t(m)$ and $\sigma_t(m',m)$ by $\Pr \{ Y_1^\tau \}$ which is equivalent to $\lambda_t(0)$, which is available from the decoder.

Alternatively we can normalize $\lambda_t(m)$ and $\sigma_t(m',m)$ to add up to 1 to obtain the same result.

\Rightarrow Derive a method for obtaining the probabilities $\lambda_t(m)$ and $\sigma_t(m',m)$.

Define the probability functions:

$$\begin{aligned} \alpha_t(m) &= \Pr \{ S_t = m ; Y_1^t \} \\ \beta_t(m) &= \Pr \{ Y_{t+1}^\tau \mid S_t = m \} \end{aligned}$$

$$\gamma_t(m',m) = \Pr \{ S_t = m ; Y_t | S_{t-1} = m' \}$$

Now $\lambda_t(m) = \Pr \{ S_t = m ; Y_1^\tau \}$

$$= \Pr \{ S_t = m ; Y_1^t \} \cdot \Pr \{ Y_{t+1}^\tau | S_t = m ; Y_1^t \}$$

where

$$\Pr \{ Y_{t+1}^\tau | S_t = m ; Y_1^t \} = \Pr \{ S_t = m ; Y_1^t ; Y_{t+1}^\tau \} / \Pr \{ S_t = m ; Y_1^t \}$$

$$= \Pr \{ S_t = m ; Y_1^\tau \} / \Pr \{ S_t = m ; Y_1^t \}$$

and $\alpha_t(m) = \Pr \{ S_t = m ; Y_1^t \}$

therefore,

$$\begin{aligned} \lambda_t(m) &= \alpha_t(m) \cdot \Pr \{ Y_{t+1}^\tau | S_t = m ; Y_1^t \} \\ &= \alpha_t(m) \cdot \Pr \{ Y_{t+1}^\tau | S_t = m \} \end{aligned}$$

Markov property: if S_t is known, events after time t do not depend on Y_1^t .

Thus $\lambda_t(m) = \alpha_t(m) \cdot \beta_t(m)$ (3)

since $\beta_t(m) = \Pr \{ Y_{t+1}^\tau | S_t = m \}$.

In a similar manner,

$$\begin{aligned} \sigma_t(m',m) &= \Pr \{ S_{t-1} = m' ; S_t = m ; Y_1^\tau \} \\ &= \Pr \{ S_{t-1} = m' ; Y_1^{t-1} \} \cdot \Pr \{ S_t = m ; Y_t | S_{t-1} = m' \} \cdot \Pr \{ Y_{t+1}^\tau | S_t = m \} \end{aligned}$$

where

$$\begin{aligned} \Pr \{ S_t = m ; Y_t | S_{t-1} = m' \} &= \Pr \{ S_t = m ; Y_t | S_{t-1} = m' ; Y_1^{t-1} \} \\ &= \Pr \{ S_t = m ; Y_t ; S_{t-1} = m' ; Y_1^{t-1} \} / \Pr \{ S_{t-1} = m' ; Y_1^{t-1} \} \end{aligned}$$

$$\begin{aligned} \Pr \{ Y_{t+1}^\tau | S_t = m \} &= \Pr \{ Y_{t+1}^\tau | S_t = m ; Y_t ; S_{t-1} = m' ; Y_1^{t-1} \} \\ &= \Pr \{ S_t = m ; Y_t ; S_{t-1} = m' ; Y_1^{t-1} ; Y_{t+1}^\tau \} / \Pr \{ S_t = m ; Y_t ; S_{t-1} = m' ; Y_1^{t-1} \} \end{aligned}$$

the equations follow from the basic probability theory and Markov property.

Thus

$$\begin{aligned}\sigma_t(m',m) &= \Pr \{ S_{t-1} = m' ; Y_1^{t-1} \} . \Pr \{ S_t = m ; Y_t | S_{t-1} = m' \} . \Pr \{ Y_{t+1}^\tau | S_t = m \} \\ &= \alpha_{t-1}(m') . \gamma_t(m',m) . \beta_t(m)\end{aligned}\quad (4)$$

Now for $t = 1, 2, \dots, \tau$

$$\begin{aligned}\mathbf{a}_t(m) &= \sum_{m'=0}^{M-1} \Pr\{S_{t-1} = m'; S_t = m; Y_1^t\} \\ &= \sum_{m'=0}^{M-1} \Pr\{S_{t-1} = m'; Y_1^{t-1}\} . \Pr\{S_t = m; Y_t | S_{t-1}=m'; Y_1^{t-1}\} \\ &= \sum_{m'=0}^{M-1} \Pr\{S_{t-1} = m'; Y_1^{t-1}\} . \Pr\{S_t = m; Y_t | S_{t-1}=m'\} \\ &= \sum_{m'=0}^{M-1} \mathbf{a}_{t-1}(m') . \mathbf{g}(m',m)\end{aligned}\quad (5)$$

For $t=0$ we have the boundary conditions

$$\alpha_0(0) = 1, \quad \text{and } \alpha_0(m) = 0, \text{ for } m \neq 0 \quad (6).$$

Similarly, for $t=1, 2, 3, \dots, \tau-1$

$$\begin{aligned}\mathbf{b}_t(m) &= \sum_{m'=0}^{M-1} \Pr\{S_{t+1} = m'; Y_{t+1}^t | S_t = m\} \\ &= \sum_{m'=0}^{M-1} \Pr\{S_{t+1} = m'; Y_{t+1} | S_t = m\} . \Pr\{Y_{t+2}^t | S_{t+1} = m'\} \\ &= \sum_{m'=0}^{M-1} \mathbf{b}_{t+1}(m') . \mathbf{g}_{+1}(m, m')\end{aligned}\quad (7)$$

The appropriate boundary conditions are

$$\beta_\tau(0) = 1, \text{ and } \beta_\tau(m) = 0, \text{ for } m \neq 0. \quad (8)$$

(5) and (7) show that $\alpha_t(m)$ and $\beta_\tau(m)$ are recursively obtainable.

Now,

$$\begin{aligned} \gamma_t(m', m) &= \sum_X \Pr\{S_t = m \mid S_{t-1} = m'\} \cdot \Pr(X_t = X \mid S_{t-1} = m', S_t = m) \cdot \Pr\{Y_t \mid X\} \\ &= \sum_X p_t(m \mid m') \cdot q_t(X \mid m', m) \cdot R(Y_t \mid X) \end{aligned} \quad (9)$$

where the summation in (9) is over all possible output symbols X .

Now the operation of the decoder for computing $\lambda_t(m)$ and $\sigma_t(m', m)$.

1. $\alpha_0(m)$ and $\beta_\tau(m)$, $m=0,1,\dots,M-1$ are initialized according to (6) and (8).
2. As soon as Y_t is received, the decoder computes $\gamma_t(m', m)$ using (9) and $\alpha_t(m)$ using (5). The obtained values for $\alpha_t(m)$ are stored for all t and m .
3. After the complete sequence Y_1^τ is received the decoder recursively computes $\beta_t(m)$ using (7). When the $\beta_t(m)$ have been computed, they can be multiplied by the appropriate $\alpha_t(m)$ and $\gamma_t(m', m)$ to obtain $\lambda_t(m)$ and $\sigma_t(m', m)$ using (3) and (4).

We can obtain the probability of any event that is a function of the states by summing the appropriate $\lambda_t(m)$; likewise, the $\sigma_t(m', m)$ can be used to obtain the probability of any event which is a function of transitions.

Summary:

- Calculation of α - “forward recursion”
- Calculation of β - “backward recursion”
- Probability of input information – by summing $\lambda_t(m)$
- Probability of coded information – by summing $\sigma_t(m',m)$

Application to Convolutional Codes

A binary k_0/n_0 Convolutional encoder. The input to the encoder at time t is the block $I_t = (i_{(t)}^{(1)}, i_{(t)}^{(2)}, \dots, i_{(t)}^{(k_0)})$ and the corresponding output is $X_t = (x_{(t)}^{(1)}, x_{(t)}^{(2)}, \dots, x_{(t)}^{(n_0)})$.

The state of the encoder (contents of the register, v most recent input blocks)

$$S_t = (s_{(t)}^{(1)}, s_{(t)}^{(2)}, \dots, s_{(t)}^{(k_0 v)}) = (I_t, I_{t-1}, \dots, I_{t-v+1}).$$

The encoder starts in state $S_0 = \mathbf{0}$. Information sequence I_1^T is the input to the encoder, followed by v blocks of all-zero inputs ,i.e., $I_{T+1}^\tau = \mathbf{0}, \mathbf{0}, \dots, \mathbf{0}$ where $\tau = T+v$, so that $S_\tau = \mathbf{0}$.

The transition probabilities $p_t(m|m')$ of the trellis are governed by the input statistics. Generally we assume all input sequences equally likely for $t \leq T$, and since there are 2^{k_0} possible transitions out of each state , $p_t(m|m') = 2^{-k_0}$ for each of these transitions.

For $t > T$, only one transition is possible out of each state, and this has probability 1.

The output X_t is a deterministic function of the transition so that, there is a 0-1 probability distribution $q_t(X|m',m)$ over the alphabet of binary n -tuples. For the time invariant codes $q_t(\cdot|\cdot)$ is independent of t .

If the output sequence is sent over a **memoryless** channel (DMC, AWGN) or where the output is conditionally independent of other outputs given the current input (ISI in AWGN) with transition probabilities $r(\cdot|\cdot)$, the derived block transition probabilities are

$$R(Y_t|X_t) = \prod_{j=1}^{n_0} r(y^{(j)} | x_t^{(j)})$$

Where $Y_t = (y_t^{(1)}, y_t^{(2)}, \dots, y_t^{(n_0)})$ is the block received by the receiver at time t .

To minimize the symbol probability of error we must determine the most likely input digits $i_t^{(j)}$ from the received sequence Y_1^τ .

Let $A_t^{(j)}$ be the set of states S_t such that $s_t^{(j)} = 0$. Note that $A_t^{(j)}$ is not dependent on t . Then we have

$$s_t^{(j)} = i_t^{(j)}, \quad j=1,2,\dots,k_0$$

which implies
$$\Pr \{ i_t^{(j)} = 0 ; Y_1^\tau \} = \sum_{s_t \in A_t^{(j)}} I_t(m)$$

Normalizing by $\Pr\{Y_1^\tau\} = \lambda_\tau(0)$

$$\Pr \{ i_t^{(j)} = 0 | Y_1^\tau \} = \frac{1}{I_t(0)} \sum_{s_t \in A_t^{(j)}} I_t(m)$$

We decode $i_t^{(j)} = 0$ if $\Pr \{ i_t^{(j)} = 0 | Y_1^\tau \} \geq 0.5$, otherwise $i_t^{(j)} = 1$.

Another equivalent way would be to compute the log likelihood ratio (LLR)

$$LLR = \text{Log} \frac{\Pr\{i_t^{(j)} = 0 | Y_1^t\}}{\Pr\{i_t^{(j)} = 1 | Y_1^t\}} = \text{Log} \frac{\sum_{s_t \in A_t^{(j)}, m} I_t(m)}{\sum_{s_t \in A_t^{(j)c}, m'} I_t(m')} \geq 0 \text{ then } i_t^{(j)} = 0.$$

Sometimes, it is of interest to determine the APP of the encoder output digits, i.e., $\Pr\{x_t^{(j)}=0 | Y_1^\tau\}$.

Let $B_t^{(j)}$ be the set of transitions $S_{t-1}=m'$ to $S_t=m$ such that the j th output digit $x_t^{(j)}$ on that transition is 0. $B_t^{(j)}$ is independent of t for time invariant codes. Then,

$$\Pr\{x_t^{(j)}=0; Y_1^\tau\} = \sum_{(m',m) \in B_t^{(j)}} \mathbf{s}_t(m',m)$$

Which can be normalized to give $\Pr\{x_t^{(j)}=0 | Y_1^\tau\}$.

Therefore we can state the following.

“We can obtain the probability of any event that is a function of the states, e.g., probability of input symbols, by summing the appropriate $\mathbf{I}_t(m)$; likewise, the $\mathbf{s}_t(m',m)$ can be used to obtain the probability of any event which is a function of transitions, e.g. probability of output symbols”.

Elements of Turbo (Iterative) Decoding

A. Analysis of Log Likelihood Ratio (LLR) of Systematic Convolutional Codes (in AWGN, Rate $(1/n_0)$ type)

The transition probability is given by

$$p(y_t^{(j)} | x_t^{(j)}) = \frac{1}{\sqrt{2ps^2}} \exp\left\{-\frac{1}{2s^2} (y_t^{(j)} - x_t^{(j)})^2\right\}$$

Now the

$$\text{LLR} = \text{Log} \frac{\Pr\{i_t = 0 | Y_1^t\}}{\Pr\{i_t = 1 | Y_1^t\}} = \text{Log} \frac{\sum_{s_t \in A_t, m} \mathbf{I}_t(m)}{\sum_{s_t \in A_t^c, m'} \mathbf{I}_t(m')}$$

$$= \text{Log} \frac{\sum_{s_t \in A_t, m, i_t=0} \mathbf{I}_t(m)}{\sum_{s_t \in A_t^c, m', i_t=1} \mathbf{I}_t(m')}$$

Using $\lambda_t(m) = \alpha_t(m)\beta_t(m)$ and $\alpha_t(m) = \sum_{m'=0}^{M-1} \mathbf{a}_{t-1}(m') \cdot \mathbf{g}(m', m)$

$$\text{LLR} = \text{Log} \frac{\sum_{s_t \in A_t, m, i_t=0} \sum_{m''=0}^{M-1} \mathbf{a}_{t-1}(m'') \cdot \mathbf{g}(m'', m) \cdot \mathbf{b}_t(m)}{\sum_{s_t \in A_t^c, m', i_t=1} \sum_{\hat{m}=0}^{M-1} \mathbf{a}_{t-1}(\hat{m}) \cdot \mathbf{g}(\hat{m}, m') \cdot \mathbf{b}_t(m')}$$

Since the encoder is systematic, i.e., $i_t = 0 \Rightarrow x_t^{(1)} = -1$ for those transitions

$$\mathbf{R}(Y_t | X_t) = p(y_t^{(1)} | x_t^{(1)} = -1) \cdot \prod_{j=2}^{n_0} p(y_t^{(j)} | x_t^{(j)}), p_t(m|m') = 1/2$$

Similarly for $i_t = 1 \Rightarrow x_t^{(1)} = 1$ and for them

$$\mathbf{R}(Y_t | X_t) = p(y_t^{(1)} | x_t^{(1)} = 1) \cdot \prod_{j=2}^{n_0} p(y_t^{(j)} | x_t^{(j)}), p_t(m|m') = 1/2$$

Therefore,

$$\text{LLR} = \text{Log} \frac{p(y_t^{(1)} | x_t^{(1)} = -1)}{p(y_t^{(1)} | x_t^{(1)} = 1)} +$$

$$\text{Log} \frac{\sum_{s_t \in A_t, m, i_t = 0} \sum_{m''=0}^{M-1} \mathbf{a}_{t-1}(m'') \cdot \prod_{j=2}^{n_0} p(y_t^{(j)} | x_t^{(j)}) \cdot \mathbf{b}_t(m)}{\sum_{s_t \in A_t^c, m', i_t = 1} \sum_{\hat{m}=0}^{M-1} \mathbf{a}_{t-1}(\hat{m}) \cdot \prod_{k=2}^{n_0} p(y_t^{(k)} | x_t^{(k)}) \cdot \mathbf{b}_t(m')}$$

The second term is a function of the redundant information introduced by the encoder. In general it has the same sign as x_t . This quantity represents the “**extrinsic**” information supplied by the decoder and does not depend on the decoder input $x_t^{(1)}$.

This property will be used for decoding “parallel concatenated codes – **TURBO** codes”.

In the first equation since we consider states with $i_t = 0$ in the numerator and the states with $i_t = 1$ respectively, we can rewrite it in the following way as well.

$$\text{LLR} = \text{Log} \frac{\sum_{(m'', m) i_t = 0} \mathbf{a}_{t-1}(m'') \cdot \underline{\mathbf{g}}_t(m'', m) \cdot \mathbf{b}_t(m)}{\sum_{(\hat{m}, m') i_t = 1} \mathbf{a}_{t-1}(\hat{m}) \cdot \underline{\mathbf{g}}_t(\hat{m}, m') \cdot \mathbf{b}_t(m')}$$

$$= \text{Log} \frac{p(y_t^{(1)} | x_t^{(1)} = -1)}{p(y_t^{(1)} | x_t^{(1)} = 1)} +$$

$$\text{Log} \frac{\sum_{(m'', m) i_t = 0} \mathbf{a}_{t-1}(m'') \cdot \prod_{j=2}^{n_0} p(y_t^{(j)} | x_t^{(j)}) \cdot \mathbf{b}_t(m)}{\sum_{(\hat{m}, m') i_t = 1} \mathbf{a}_{t-1}(\hat{m}) \cdot \prod_{k=2}^{n_0} p(y_t^{(k)} | x_t^{(k)}) \cdot \mathbf{b}_t(m')}$$

Thus the extrinsic part can be represented by

$$\text{Log} \frac{\sum_{(m'', m) i_t = 0} \mathbf{a}_{t-1}(m'') \cdot \boldsymbol{\xi}_{te}(m'', m) \cdot \mathbf{b}_t(m)}{\sum_{(\hat{m}, m') i_t = 1} \mathbf{a}_{t-1}(\hat{m}) \cdot \boldsymbol{\xi}_{te}(\hat{m}, m') \cdot \mathbf{b}_t(m')}$$