

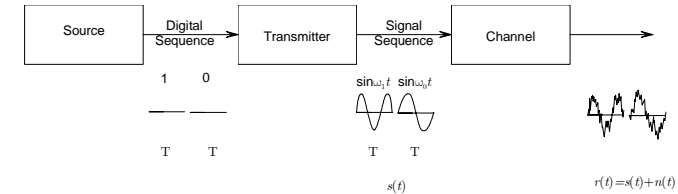
## Bayesian Hypothesis testing

Statistical Decision Theory I.  
 Simple Hypothesis testing.  
     Binary Hypothesis testing  
     Bayesian Hypothesis testing.  
 Minimax Hypothesis testing.  
 Neyman-Pearson criterion.  
 M-Hypotheses.  
 Receiver Operating Characteristics.  
 Composite Hypothesis testing.  
 Composite Hypothesis testing approaches.  
 Performance of GLRT for large data records.  
 Nuisance parameters.

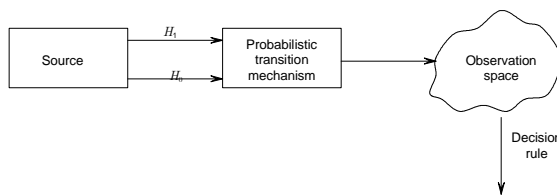
## Classical detection and estimation theory.

What is detection?

- Signal detection and estimation is the area of study that deals with the processing of information-bearing signals for the purpose of extracting information from them.



## Components of a decision theory problem.

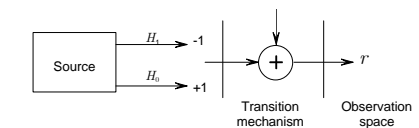
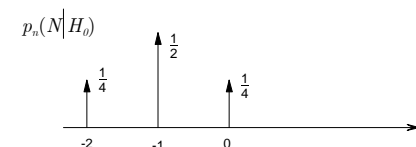
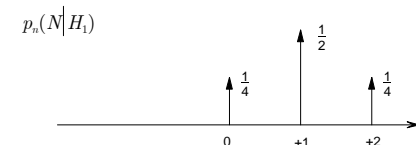
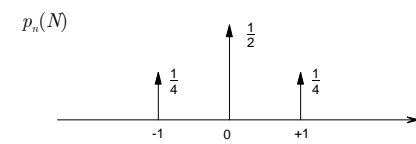


Components of a decision theory problem.

1. Source - that generates an output.
2. Probabilistic transition mechanism - a device that knows which hypothesis is true. It generates a point in the observation space accordingly to some probability law.

3. Observation space – describes all the outcomes of the transition mechanism.
4. Decision - to each point in observation space is assigned one of the hypotheses

## Example:



- When  $H_1$  is true the source generates +1.
- When  $H_0$  is true the source generates -1.
- An independent discrete random variable  $n$  whose probability density is added to the source output.

- The sum of the source output and  $n$  is observed variable  $r$ .
- Observation space has finite dimension, i.e. observation consists of a set of  $N$  numbers and can be represented as a point in  $N$  dimensional space.
- Under the two hypotheses, we have
 
$$H_1 : r = 1 + n$$

$$H_0 : r = -1 + n$$
- After observing the outcome in the observation space we shall guess which hypothesis is true.
- We use a decision rule that assigns each point to one of the hypotheses.

- Detection and estimation applications involve making inferences from observations that are distorted or corrupted in some unknown manner.

### Simple binary hypothesis testing.

- The decision problem in which each of two source outputs corresponds to a hypothesis.
- Each hypothesis maps into a point in the observation space.
- We assume that the observation space is a set of  $N$  observations:
 
$$r_1, r_2, \dots, r_N.$$
- Each set can be represented as a vector  $\mathbf{r}$ :

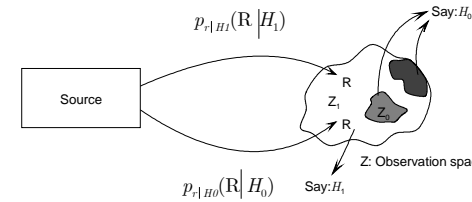
$$\mathbf{r} \triangleq \begin{bmatrix} r_1 \\ r_2 \\ \vdots \\ r_N \end{bmatrix}$$

- The probabilistic transition mechanism generates points in accord with the two known conditional densities  $p_{\mathbf{r}|H_1}(\mathbf{R} | H_1), p_{\mathbf{r}|H_0}(\mathbf{R} | H_0)$ .
- The objective is to use this information to develop a decision rule.

**Decision criteria.**

- In the binary hypothesis problem either  $H_0$  or  $H_1$  is true.
- We are seeking decision rules for making a choice.
- Each time the experiment is conducted one of four things can happen:
  1.  $H_0$  true; choose  $H_0 \rightarrow$  correct
  2.  $H_0$  true; choose  $H_1$
  3.  $H_1$  true; choose  $H_1 \rightarrow$  correct
  4.  $H_1$  true; choose  $H_0$
- The purpose of a decision criterion is to attach some relative importance to the four possible courses of action.
- The method for processing the received data depends on the decision criterion we select.

**Bayesian criterion.**



Source generates two outputs with given (*a priori*) probabilities  $P_1, P_0$ . These represent the observer information before the experiment is conducted.

- The cost is assigned to each course of actions.  $C_{00}, C_{10}, C_{01}, C_{11}$ .
- Each time the experiment is conducted a certain cost will be incurred.
- The decision rule is designed so that on the average the cost will be as small as possible.
- Two probabilities are averaged over: the *a priori* probability and probability that a particular course of action will be taken.

- The expected value of the cost is

$$\begin{aligned} \mathbf{R} = & C_{00}P_0 \Pr(\text{ say } H_0 | H_0 \text{ is true}) \\ & + C_{10}P_0 \Pr(\text{ say } H_1 | H_0 \text{ is true}) \\ & + C_{11}P_1 \Pr(\text{ say } H_1 | H_1 \text{ is true}) \\ & + C_{01}P_1 \Pr(\text{ say } H_0 | H_1 \text{ is true}) \end{aligned}$$

- The binary observation rule divides the total observation space  $Z$  into two parts:  $Z_0, Z_1$ .
- Each point in observation space is assigned to one of these sets.
- The expression of the risk in terms of transition probabilities and the decision regions:

$$\begin{aligned} \mathbf{R} = & C_{00}P_0 \int_{Z_0} p_{r|H_0}(\mathbf{R} | H_0) d\mathbf{R} + C_{10}P_0 \int_{Z_1} p_{r|H_0}(\mathbf{R} | H_0) d\mathbf{R} \\ & + C_{11}P_1 \int_{Z_1} p_{r|H_1}(\mathbf{R} | H_1) d\mathbf{R} + C_{01}P_1 \int_{Z_0} p_{r|H_1}(\mathbf{R} | H_1) d\mathbf{R} \end{aligned}$$

- $Z_0, Z_1$  cover the observation space (the integrals integrate to one).
- We assume that the cost of a wrong decision is higher than the cost of a correct decision.
 
$$C_{10} > C_{00}$$

$$C_{01} > C_{11}$$
- For Bayesian test the regions  $Z_0$  and  $Z_1$  are chosen such that the risk will be minimized.

- We assume that the decision is to be made for each point in observation space. ( $Z = Z_0 + Z_1$ )

- The decision regions are defined by the statement:

$$\mathbf{R} = C_{00}P_0 \int_{Z_0} p_{\mathbf{r}|H_0}(\mathbf{R} | H_0) d\mathbf{R} + C_{10}P_0 \int_{Z-Z_0} p_{\mathbf{r}|H_0}(\mathbf{R} | H_0) d\mathbf{R} \\ + C_{11}P_1 \int_{Z-Z_0} p_{\mathbf{r}|H_1}(\mathbf{R} | H_1) d\mathbf{R} + C_{01}P_1 \int_{Z_0} p_{\mathbf{r}|H_1}(\mathbf{R} | H_1) d\mathbf{R}$$

Observing that

$$\int_Z p_{\mathbf{r}|H_0}(\mathbf{R} | H_0) d\mathbf{R} = \int_Z p_{\mathbf{r}|H_1}(\mathbf{R} | H_1) d\mathbf{R} = 1$$

$$\mathbf{R} = P_0C_{10} + P_1C_{11} + \int_{Z_0} \left\{ \begin{array}{l} P_1(C_{01} - C_{11})p_{\mathbf{r}|H_1}(\mathbf{R} | H_1) \\ -P_0(C_{10} - C_{00})p_{\mathbf{r}|H_0}(\mathbf{R} | H_0) \end{array} \right\} d\mathbf{R}$$

- The integral represents the cost controlled by those points  $\mathbf{R}$  that we assign to  $Z_0$ .
- The value of  $\mathbf{R}$  where the second term is larger than the first contribute to the negative amount to the integral and should be included in  $Z_0$ .
- The value of  $\mathbf{R}$  where two terms are equal has no effect.
- The decision regions are defined by the statement:  
If  $P_1(C_{01} - C_{11})p_{\mathbf{r}|H_1}(\mathbf{R} | H_1) \geq P_0(C_{10} - C_{00})p_{\mathbf{r}|H_0}(\mathbf{R} | H_0)$ , assign  $\mathbf{R}$  to  $Z_1$  and say that  $H_1$  is true. Otherwise assign  $\mathbf{R}$  to  $Z_0$  and say that  $H_0$  is true.

- This may be expressed as:

$$\frac{p_{\mathbf{r}|H_1}(\mathbf{R} | H_1)}{p_{\mathbf{r}|H_0}(\mathbf{R} | H_0)} \underset{H_1}{\overset{H_0}{\gtrless}} \frac{P_0(C_{10} - C_{00})}{P_1(C_{01} - C_{11})}$$

- $\Lambda(\mathbf{R}) = \frac{p_{\mathbf{r}|H_1}(\mathbf{R} | H_1)}{p_{\mathbf{r}|H_0}(\mathbf{R} | H_0)}$  is called likelihood ratio.
- Regardless of the dimension of  $\mathbf{R}$ ,  $\Lambda(\mathbf{R})$  is one-dimensional variable.
- Data processing is involved in computing  $\Lambda(\mathbf{R})$  and is not affected by the prior probabilities and cost assignments.
- The quantity  $\eta \triangleq \frac{P_0(C_{10} - C_{00})}{P_1(C_{01} - C_{11})}$  is the threshold of the test.

- The  $\eta$  can be left as a variable threshold and may be changed if our a priori knowledge or costs are changed.
- Bayes criterion has led us to a Likelihood Ratio Test (LRT)

$$\Lambda(\mathbf{R}) \underset{H_1}{\overset{H_0}{\gtrless}} \eta$$

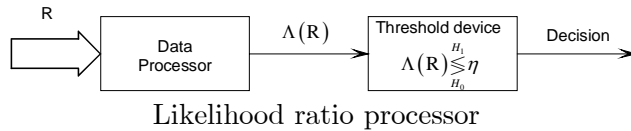
- An equivalent test is  $\ln \Lambda(\mathbf{R}) \underset{H_1}{\overset{H_0}{\gtrless}} \ln \eta$

**Summary of the Bayesian test:**

- The Bayesian test can be conducted simply by calculating the likelihood ratio  $\Lambda(\mathbf{R})$  and comparing it to the threshold.

Test design:

- Assign a-priori probabilities to the source outputs.
- Assign costs for each action.
- Assume distribution for  $p_{\mathbf{r}|H_1}(\mathbf{R} | H_1), p_{\mathbf{r}|H_0}(\mathbf{R} | H_0)$ .
- calculate and simplify the  $\Lambda(\mathbf{R})$



**Special case.**

$$C_{00} = C_{11} = 0 \quad C_{01} = C_{10} = 1$$

$$\mathbf{R} = P_0 \int_{Z-Z_0} p_{\mathbf{r}|H_0}(\mathbf{R} | H_0) d\mathbf{R} + P_1 \int_{Z_0} p_{\mathbf{r}|H_1}(\mathbf{R} | H_1) d\mathbf{R}$$

$$\ln \Lambda(\mathbf{R}) \underset{H_1}{\overset{H_0}{\approx}} \ln \frac{P_0}{P_1} = \ln P_0 - \ln(1 - P_1)$$

- When the two hypotheses are equally likely, the threshold is zero.

**Sufficient statistics.**

- Sufficient statistics is a function  $T$  that transfers the initial data set to the new data set  $T(\mathbf{R})$  that still contains all necessary information contained in  $\mathbf{R}$  regarding the problem under investigation.
- The set that contains a minimal amount of elements is called minimal sufficient statistics.
- When making a decision knowing the value of the sufficient statistic is just as good as knowing  $\mathbf{R}$ .

**The integrals in the Bayes test.**

False alarm:

$$P_F = \int_{Z_1} p_{\mathbf{r}|H_0}(\mathbf{R} | H_0) d\mathbf{R}.$$

We say that target is present when it is not.

Probability of detection:

$$P_D = \int_{Z_1} p_{\mathbf{r}|H_1}(\mathbf{R} | H_1) d\mathbf{R}.$$

Probability of miss:

$$P_M = \int_{Z_0} p_{\mathbf{r}|H_1}(\mathbf{R} | H_1) d\mathbf{R}.$$

We say target is absent when it is present.

**Special case: the prior probabilities unknown.**

**Minimax test.**

$$\mathbf{R} = C_{00}P_0 \int_{Z_0} p_{r|H_0}(\mathbf{R} | H_0) d\mathbf{R} + C_{10}P_0 \int_{Z-Z_0} p_{r|H_0}(\mathbf{R} | H_0) d\mathbf{R}$$

$$+ C_{11}P_1 \int_{Z-Z_0} p_{r|H_1}(\mathbf{R} | H_1) d\mathbf{R} + C_{01}P_1 \int_{Z_0} p_{r|H_1}(\mathbf{R} | H_1) d\mathbf{R}$$

- If the regions  $Z_0$  and  $Z_1$  fixed the integrals are determined.

$$\mathbf{R} = P_0 C_{10} + P_1 C_{11} + P_1 (C_{01} - C_{11}) P_M - P_0 (C_{10} - C_{00}) (1 - P_F)$$

$$P_0 = 1 - P_1$$

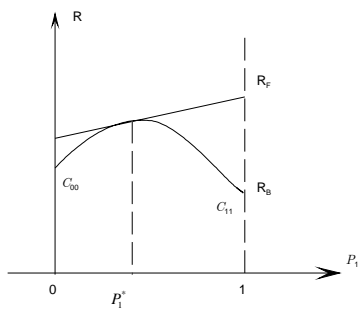
- The Bayesian risk will be function of  $P_1$ .

$$\mathbf{R}(P_1) = C_{00}(1 - P_F) + C_{10}P_F$$

$$+ P_1 [(C_{11} - C_{00}) + (C_{01} - C_{11})P_M - (C_{10} - C_{00})P_F]$$

- Bayesian test can be found if all the costs and a priori probabilities are known.
- If we know all the probabilities we can calculate the Bayesian cost.
- Assume that we do not know  $P_1$  and just assume a certain one  $P_1^*$  and design a corresponding test.
- If  $P_1$  changes the regions  $Z_0, Z_1$  changes and with these also  $P_F$  and  $P_D$ .

- The test is designed for  $P_1^*$  but the actual a priori probability is  $P_1$ .



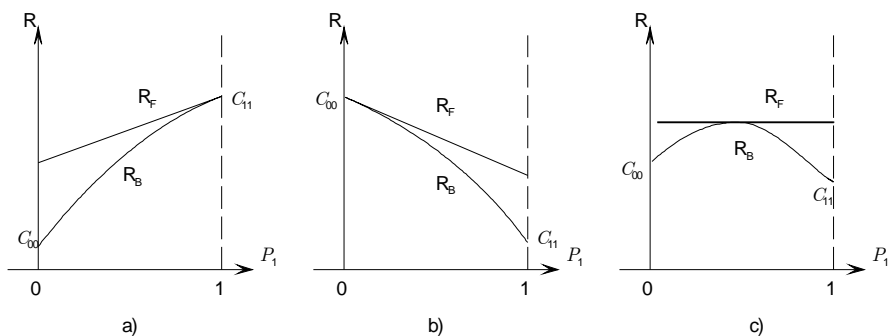
A function of  $P_1$  if  $P_1^*$  is fixed

- By assuming  $P_1^*$  we fix  $P_F$  and  $P_D$ .
  - Cost for different  $P_1$  is given by a function  $\mathbf{R}(P_1^*, P_1)$ .
  - Because the threshold  $\eta$  is fixed the cost  $\mathbf{R}(P_1^*, P_1)$  is a linear function of  $P_1$ .
  - Bayesian test minimizes the risk for  $P_1^*$ .
- For other values of  $P_1$
- $$\mathbf{R}(P_1^*, P_1) \geq \mathbf{R}(P_1)$$
- $\mathbf{R}(P_1)$  is strictly concave. (If  $\Lambda(\mathbf{R})$  is a continuous random variable with strictly

monotonic probability distribution function, the change of  $\eta$  always change the risk.)

**Minimax test.**

- The Bayesian test designed to minimize the maximum possible risk is called a minimax test.
- $P_1$  is chosen to maximize our risk  $\mathbf{R}(P_1^*, P_1)$ .
- To minimize the maximum risk we select the  $P_1^*$  for which  $\mathbf{R}(P_1)$  is maximum.
- If the maximum occurs inside the interval  $[0,1]$ , the  $\mathbf{R}(P_1^*, P_1)$  will become a horizontal line. Coefficient of  $P_1$  must be zero.
- $(C_{11} - C_{00}) + (C_{01} - C_{11})P_M - (C_{10} - C_{00})P_F = 0$  This equation is the minimax equation.



Risk curves: maximum value of  $\mathbf{R}$  at a)  $P_1 = 1$  b)  $P_1 = 0$  c)  $0 \leq P_1 \leq 1$

**Special case.**

Cost function is

$$C_{00} = C_{11} = 0$$

$$C_{01} = C_M, C_{10} = C_F.$$

The risk is

$$\mathbf{R}(P_1) = C_F P_F + P_1 (C_M P_M - C_F P_F) = P_0 C_F P_F + P_1 C_M P_M.$$

The minimax equation is

$$C_M P_M = C_F P_F.$$

**Neyman-Pearson test.**

- Often it is difficult to assign realistic costs of a priori probabilities. This can be bypassed if to work with the conditional probabilities  $P_F$  and  $P_D$ .
- We have two conflicting objectives to make  $P_F$  as small as possible and  $P_D$  as large as possible.

**Neyman-Pearson criterion.**

Constrain  $P_F = \alpha' \leq \alpha$  and design a test to maximize  $P_D$  (or minimize  $P_M$ ) under this constraint.

- The solution can be obtained by using Lagrange multipliers.

$$F = P_M + \lambda [P_F - \alpha']$$

$$F = \int_{Z_0} p_{r|H_1}(\mathbf{R} | H_1) d\mathbf{R} + \lambda \left[ \int_{Z-Z_0} p_{r|H_0}(\mathbf{R} | H_0) d\mathbf{R} - \alpha' \right]$$

- If  $P_F = \alpha$ , minimizing  $F$  minimizes  $P_M$ .

$$F = \lambda (1 - \alpha') + \int_{Z_0} [p_{r|H_1}(\mathbf{R} | H_1) - \lambda p_{r|H_0}(\mathbf{R} | H_0)] d\mathbf{R}$$

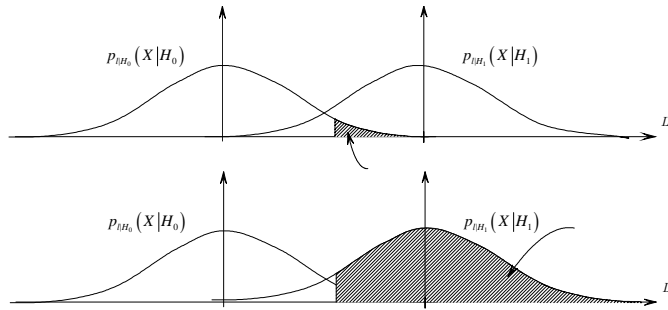
- For any positive value of  $\lambda$  an LRT will minimize  $F$ .
- $F$  is minimized if we assign a point  $\mathbf{R}$  to  $Z_0$  only when the term in the bracket is negative.

- If  $\frac{p_{r|H_1}(\mathbf{R} | H_1)}{p_{r|H_0}(\mathbf{R} | H_0)} < \lambda$  assign point to  $Z_0$  (say  $H_0$ )

- F is minimized by the likelihood ratio test.  $\Lambda(\mathbf{R}) \underset{H_0}{\overset{H_1}{\leq}} \eta$
- To satisfy the constraint  $\lambda$  is selected so that  $P_F = \alpha'$ .

$$P_F = \int_{\lambda}^{\infty} p_{\Lambda|H_0}(\Lambda | H_0) d\Lambda = \alpha'$$

- Value of  $\lambda$  will be nonnegative because  $p_{\Lambda|H_0}(\Lambda | H_0)$  will be zero for negative values of  $\lambda$ .



$$p_{r_i|H_1}(R_i | H_1) = p_{n_i}(R_i - m) = \frac{1}{\sqrt{2\pi}\sigma} \exp\left(-\frac{(R_i - m)^2}{2\sigma^2}\right)$$

$$p_{r_i|H_0}(R_i | H_0) = p_{n_i}(R_i) = \frac{1}{\sqrt{2\pi}\sigma} \exp\left(-\frac{R_i^2}{2\sigma^2}\right)$$

- The joint probability of  $N$  samples is:

$$p_{\mathbf{r}|H_1}(\mathbf{R} | H_1) = \prod_{i=1}^N \frac{1}{\sqrt{2\pi}\sigma} \exp\left(-\frac{(R_i - m)^2}{2\sigma^2}\right)$$

$$p_{\mathbf{r}|H_0}(\mathbf{R} | H_0) = \prod_{i=1}^N \frac{1}{\sqrt{2\pi}\sigma} \exp\left(-\frac{R_i^2}{2\sigma^2}\right)$$

### Example.

We assume that under  $H_1$  the source output is a constant voltage  $m$ . Under  $H_0$  the source output is zero. Voltage is corrupted by an additive noise. The output is sampled with  $N$  samples for each second. Each noise sample is a i.i.d. zero mean Gaussian random variable with variance  $\sigma^2$ .

$$H_1 : r_i = m + n_i, \quad i = 1, 2, \dots, N$$

$$H_0 : r_i = n_i, \quad i = 1, 2, \dots, N$$

$$p_{n_i}(X) = \frac{1}{\sqrt{2\pi}\sigma} \exp\left(-\frac{X^2}{2\sigma^2}\right)$$

The probability density of  $r_i$  under each hypothesis is:

- The likelihood ratio is

$$\Lambda(\mathbf{R}) = \frac{\prod_{i=1}^N \frac{1}{\sqrt{2\pi}\sigma} \exp\left(-\frac{(R_i - m)^2}{2\sigma^2}\right)}{\prod_{i=1}^N \frac{1}{\sqrt{2\pi}\sigma} \exp\left(-\frac{R_i^2}{2\sigma^2}\right)}$$

- After cancelling common terms and taking logarithm:

$$\ln \Lambda(\mathbf{R}) = \frac{m}{\sigma^2} \sum_{i=1}^N R_i - \frac{Nm^2}{2\sigma^2}.$$

- Likelihood ratio test is:



$$\frac{m}{\sigma^2} \sum_{i=1}^N R_i - \frac{Nm^2}{2\sigma^2} \underset{H_0}{\overset{H_1}{\leq}} \ln \eta,$$

$$\sum_{i=1}^N R_i \underset{H_0}{\overset{H_1}{\leq}} \frac{\sigma^2}{m} \ln \eta + \frac{Nm}{2} \triangleq \gamma.$$

- We use  $d \triangleq \frac{\sqrt{Nm}}{\sigma}$  for normalisation.

$$l = \frac{1}{\sqrt{N}\sigma} \sum_{i=1}^N R_i \underset{H_0}{\overset{H_1}{\leq}} \frac{\sigma}{\sqrt{Nm}} \ln \eta + \frac{\sqrt{Nm}}{2\sigma}$$

- Under  $H_0$   $l$  is obtained by adding  $N$  independent zero mean Gaussian variables with variance  $\sigma^2$  and then dividing by  $\sqrt{N}\sigma$ . Therefore  $l$  is  $N(0,1)$ .

- Under  $H_1$   $l$  is  $N\left(\frac{\sqrt{Nm}}{\sigma}, 1\right)$ .

$$P_F = \int_{(\log \eta)/d + d/2}^{\infty} \frac{1}{\sqrt{2}} \exp\left(-\frac{x^2}{2}\right) dx = \text{erfc}\left(\frac{\ln \eta}{d} + \frac{d}{2}\right)$$

where  $d \triangleq \frac{\sqrt{Nm}}{\sigma}$  is the distance between the means of the two densities.

$$P_D = \int_{(\log \eta)/d + d/2}^{\infty} \frac{1}{\sqrt{2}} \exp\left(-\frac{(x-d)^2}{2}\right) dx$$

$$= \int_{(\log \eta)/d - d/2}^{\infty} \frac{1}{\sqrt{2}} \exp\left(-\frac{(y)^2}{2}\right) dy = \text{erfc}\left(\frac{\log \eta}{d} - \frac{d}{2}\right)$$

- In the communication systems a special case is important

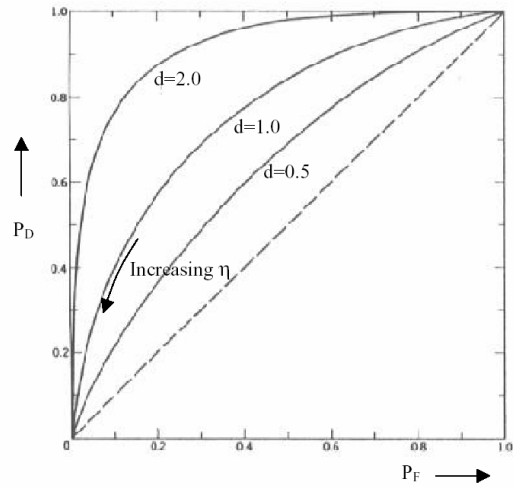
$$\Pr(\varepsilon) \triangleq P_0 P_F + P_1 P_M.$$

- If  $P_0 = P_1$  the threshold is one and  $\Pr(\varepsilon) \triangleq \frac{1}{2}(P_F + P_M)$ .

### **Receiver Operating Characteristics (ROC).**

- For a Neyman-Pearson test the values of  $P_F$  and  $P_D$  completely specify the test performance.
- $P_D$  depends on  $P_F$ . The function of  $P_D(P_F)$  is defined as the Receiver Operating Characteristic (ROC).
- The Receiver Operating Characteristic (ROC) completely describes the performance of the test as a function of the parameters of interest.

**Example.**



Receiver Operating Characteristic (ROC)

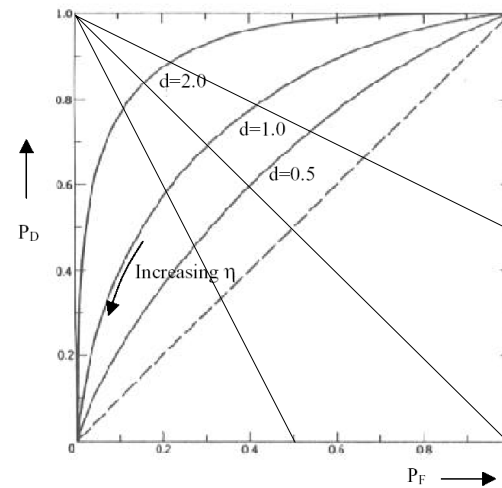
**Properties of ROC.**

- All continuous likelihood tests have ROC's that are concave downward.
- All continuous likelihood ratio tests have ROC's that are above the  $P_D = P_F$ .
- The slope of a curve in a ROC at a particular point is equal to the value of the threshold  $\eta$  required to achieve the  $P_F$  and  $P_D$  at that point

Whenever the maximum value of the Bayes risk is interior to the interval (0,1) of the P1 axis the minimax operating point is the intersection of the line

$$(C_{11} - C_{00}) + (C_{01} - C_{11})(1 - P_D) - (C_{10} - C_{00})P_F = 0$$

and the appropriate curve on the ROC.



Receiver Operating Characteristic (ROC)

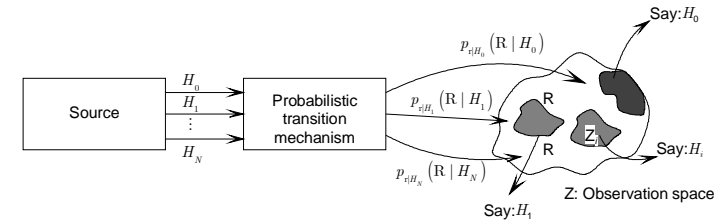
Determination of minimax operating point.

### Conclusions.

- Using either the Bayes criterion of a Neyman-Pearson criterion, we find that the optimum test is a likelihood ratio test.
- Regardless of the dimension of the observation space the optimum test consist of comparing a scalar variable  $\Lambda(\mathbf{R})$  with the threshold.
- For the binary hypothesis test the decision space is one dimensional.
- The test can be simplified by calculating the sufficient statistics.
- A complete description of the LRT performance was obtained by plotting the conditioning probabilities  $P_D$  and  $P_F$  as the threshold  $\eta$  was varied.

### M Hypotheses.

- We choose one of M hypotheses
- There are M source outputs each of which corresponds to one of M hypotheses.
- We are forced to make decisions.
- There are  $M^2$  alternatives that may occur each time the experiment is conducted.



### Bayes Criterion.

$C_{ij}$  cost of each course of actions.

$Z_i$  region in observation space where we chose  $H_i$

$P_i$  a priori probabilities.

$$\mathbf{R} = \sum_{i=0}^{M-1} \sum_{j=0}^{M-1} P_j C_{ij} \int_{Z_i} p_{\mathbf{r}|H_j}(\mathbf{R} | H_j) d\mathbf{R}$$

$\mathbf{R}$  is minimized through selecting  $Z_i$ .

### **Example M = 3.**

$$Z = Z_0 + Z_1 + Z_2$$

$$\begin{aligned} \mathbf{R} = & P_0 C_{00} \int_{Z-Z_1-Z_2} p_{\mathbf{r}|H_0}(\mathbf{R} | H_0) d\mathbf{R} + P_0 C_{10} \int_{Z_1} p_{\mathbf{r}|H_0}(\mathbf{R} | H_0) d\mathbf{R} \\ & + P_0 C_{20} \int_{Z_2} p_{\mathbf{r}|H_0}(\mathbf{R} | H_0) d\mathbf{R} + P_1 C_{11} \int_{Z-Z_0-Z_2} p_{\mathbf{r}|H_1}(\mathbf{R} | H_1) d\mathbf{R} \\ & + P_1 C_{01} \int_{Z_0} p_{\mathbf{r}|H_1}(\mathbf{R} | H_1) d\mathbf{R} + P_1 C_{21} \int_{Z_2} p_{\mathbf{r}|H_1}(\mathbf{R} | H_1) d\mathbf{R} \\ & + P_2 C_{22} \int_{Z-Z_0-Z_1} p_{\mathbf{r}|H_2}(\mathbf{R} | H_2) d\mathbf{R} + P_2 C_{02} \int_{Z_0} p_{\mathbf{r}|H_2}(\mathbf{R} | H_2) d\mathbf{R} \\ & + P_2 C_{12} \int_{Z_1} p_{\mathbf{r}|H_2}(\mathbf{R} | H_2) d\mathbf{R} \end{aligned}$$

$$\begin{aligned} \mathbf{R} &= P_0 C_{00} + P_1 C_{11} + P_2 C_{22} \\ &+ \int_{Z_0} \left[ P_2 (C_{02} - C_{22}) p_{\mathbf{r}|H_2}(\mathbf{R} | H_2) + P_1 (C_{01} - C_{11}) p_{\mathbf{r}|H_1}(\mathbf{R} | H_1) \right] d\mathbf{R} \\ &+ \int_{Z_1} \left[ P_0 (C_{10} - C_{00}) p_{\mathbf{r}|H_0}(\mathbf{R} | H_0) + P_2 (C_{12} - C_{22}) p_{\mathbf{r}|H_2}(\mathbf{R} | H_2) \right] d\mathbf{R} \\ &+ \int_{Z_2} \left[ P_0 (C_{20} - C_{00}) p_{\mathbf{r}|H_0}(\mathbf{R} | H_0) + P_1 (C_{21} - C_{11}) p_{\mathbf{r}|H_1}(\mathbf{R} | H_1) \right] d\mathbf{R} \end{aligned}$$

- $\mathbf{R}$  is minimized if we assign each  $\mathbf{R}$  to the region in which the value of the integrand is the smallest.
- Label the integrals  $I_0(\mathbf{R}), I_1(\mathbf{R}), I_2(\mathbf{R})$ .

- $I_0(\mathbf{R}) < I_1(\mathbf{R})$  and  $I_2(\mathbf{R})$ , choose  $H_0$
- $I_1(\mathbf{R}) < I_0(\mathbf{R})$  and  $I_2(\mathbf{R})$ , choose  $H_1$
- $I_2(\mathbf{R}) < I_0(\mathbf{R})$  and  $I_1(\mathbf{R})$ , choose  $H_2$

- If we use likelihood ratios

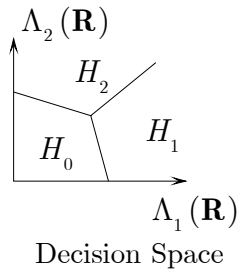
$$\Lambda_1(\mathbf{R}) \triangleq \frac{p_{\mathbf{r}|H_1}(\mathbf{R} | H_1)}{p_{\mathbf{r}|H_0}(\mathbf{R} | H_0)}$$

$$\Lambda_2(\mathbf{R}) \triangleq \frac{p_{\mathbf{r}|H_2}(\mathbf{R} | H_2)}{p_{\mathbf{r}|H_0}(\mathbf{R} | H_0)}$$

The set of decision equations is:

$$\begin{aligned} P_1 (C_{01} - C_{11}) \Lambda_1(\mathbf{R}) &\stackrel{H_0 \text{ or } H_2}{\underset{H_1 \text{ or } H_2}{\leq}} P_0 (C_{10} - C_{00}) + P_2 (C_{12} - C_{02}) \Lambda_2(\mathbf{R}) \\ P_2 (C_{02} - C_{22}) \Lambda_2(\mathbf{R}) &\stackrel{H_0 \text{ or } H_1}{\underset{H_2 \text{ or } H_1}{\leq}} P_0 (C_{20} - C_{00}) + P_1 (C_{21} - C_{01}) \Lambda_1(\mathbf{R}) \\ P_2 (C_{12} - C_{22}) \Lambda_2(\mathbf{R}) &\stackrel{H_0 \text{ or } H_1}{\underset{H_0 \text{ or } H_2}{\leq}} P_0 (C_{20} - C_{10}) + P_1 (C_{21} - C_{11}) \Lambda_1(\mathbf{R}) \end{aligned}$$

- $M$  hypotheses always lead to a decision space that has, at most,  $M - 1$  dimensions.



**Special case.** (often in communication)

$$C_{00} = C_{11} = C_{22} = 0$$

$$C_{ij} = 1, \quad i \neq j$$

$$P_1 p_{\mathbf{r}|H_1}(\mathbf{R} | H_1) \stackrel{H_1 \text{ or } H_2}{\underset{H_0 \text{ or } H_2}{\leq}} P_0 p_{\mathbf{r}|H_0}(\mathbf{R} | H_0)$$

$$P_2 p_{\mathbf{r}|H_2}(\mathbf{R} | H_2) \stackrel{H_1 \text{ or } H_2}{\underset{H_0 \text{ or } H_1}{\leq}} P_0 p_{\mathbf{r}|H_0}(\mathbf{R} | H_0)$$

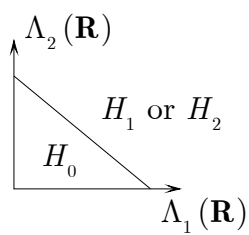
$$P_2 p_{\mathbf{r}|H_2}(\mathbf{R} | H_2) \stackrel{H_0 \text{ or } H_2}{\underset{H_0 \text{ or } H_1}{\leq}} P_0 p_{\mathbf{r}|H_1}(\mathbf{R} | H_1)$$

- Compute the posterior probabilities and choose the largest.
- Maximum a posteriori probability computer.

**Special case.**

Degeneration of hypothesis.

What happens if to combine  $H_1$  and  $H_2$ :



$$C_{12} = C_{21} = 0.$$

For simplicity

$$C_{01} = C_{10} = C_{20} = C_{02}$$

$$C_{00} = C_{11} = C_{22} = 0$$

First two equations of the test reduce to

$$P_1 \Lambda_1(\mathbf{R}) + P_2 \Lambda_2(\mathbf{R}) \underset{H_0}{\overset{H_1 \text{ or } H_2}{\leq}} P_0$$

**Dummy hypothesis.**

- Actual problem has two hypothesis  $H_1$  and  $H_2$ .
- We introduce a new one  $H_0$  with a priori probability  $P_0 = 0$

• Let

$$P_1 + P_2 = 1 \text{ and } C_{12} = C_{02}, C_{21} = C_{01}$$

- We always choose  $H_1$  or  $H_2$ . The test reduces to:

$$P_2 (C_{12} - C_{22}) \Lambda_2(\mathbf{R}) \underset{H_1}{\overset{H_2}{\leq}} P_1 (C_{21} - C_{11}) \Lambda_1(\mathbf{R})$$

- Useful if  $\frac{p_{r|H_2}(\mathbf{R} | H_2)}{p_{r|H_1}(\mathbf{R} | H_1)}$  difficult to work with but  $\Lambda_1(\mathbf{R}), \Lambda_2(\mathbf{R})$  are simple.

**Conclusions.**

1. The minimum dimension of the decision space is no more than  $M - 1$ . The boundaries of the decision regions are hyperplanes in the  $(\Lambda_1, \dots, \Lambda_{m-1})$  plane.
2. The test is simple to find but error probabilities are often difficult to compute.
3. An important test is the minimum total probability of error test. Here we compute the a posteriori probability of each test and choose the largest.